# Working Paper



# UIP Deviations in Times of Uncertainty: Not all Countries Behave Alike

Purva Gole<sup>†</sup>, Erica Perego<sup>††</sup> & Camelia Turcu<sup>‡</sup>

# Highlights

- We reconsider the role of uncertainty in explaining uncovered interest rate parity (UIP) deviations by focusing on 60 emerging and developing (EMDE) and advanced (AE) economies.
- We show that differentiating between EMDE currencies and AE currencies is crucial for understanding UIP deviations.
- In periods of uncertainty UIP deviations become wider for EMDEs and narrower for AEs.



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# Abstract

In this paper, we reconsider the role of uncertainty in explaining uncovered interest rate parity (UIP) deviations by focusing on 60 emerging and developing (EMDE) and advanced (AE) economies, over the period 1995M1--2023M3. We show that differentiating between EMDE currencies and AE currencies is crucial for understanding UIP deviations as the behaviour of excess returns differs in the two groups in periods of uncertainty: deviations become wider for EMDEs and narrow for AEs. These new results are consistent with the idea that in periods of uncertainty, global investors might change their risk preferences and move from high currency-risk investments in EMDEs towards less risky ones in AEs. This evidence holds for both the short-run and long-run UIP, and becomes stronger since the Global Financial Crisis (GFC).

# Keywords

Uncertainty, Uncovered Interest Rate Parity, Risk Premia, Emerging Countries.



F21, F30, F31, F41.

## Working Paper



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#### 1 Introduction

The theory of the uncovered interest rate parity (UIP henceforth) suggests that identical assets in two different countries should provide the same returns, once taken into account the exchange rate behaviour of the two countries' currencies. However, the empirical literature has shown—since Fama (1984)'s seminal work—that the UIP condition does not hold in reality. The literature, focusing on developed countries, has identified different explanations: the role of expectations, uncertainty, and country risk premium.

In this paper, we reconsider the role of uncertainty in explaining UIP deviations. We conduct a comprehensive study of monthly UIP deviations in periods of uncertainty by considering a large sample of 60 currencies including both EMDEs and AEs countries. Our contribution is to show that differentiating between EMDE currencies and AE currencies is crucial for understanding UIP deviations as the behaviour of excess returns differs in the two groups: in periods of uncertainty UIP deviations become wider for EMDEs and narrower for AEs.

Heightened uncertainty can increase investors' risk aversion and push investors to postpone their investment decisions and/or to require higher carry trade payoffs (excess returns) as compensation for bearing currency risk. This behaviour blurs the relationship between exchange rates and interest rate differentials leading to the so-called UIP risk premium. In this paper, we conjecture that, when arbitrage opportunity gains become more uncertain, global investors might change their risk preferences and decide to move their investments from high currency risk countries (EMDEs) towards less risky ones (AEs) determining a higher UIP risk premium for the former, and a lower for the latter.

We contribute to the literature in two directions. First, we enlarge the recent literature on UIP and uncertainty by focusing on both EMDEs and AEs. Previous studies have shown that UIP holds in periods of low uncertainty and breaks down during periods of high uncertainty (Ismailov and Rossi, 2018; Husted et al., 2018), and that uncertainty is priced as a global risk in excess returns (Berg and Mark, 2018; Ferrara and Yapi, 2022). Second, we contribute to the literature by studying the differentiated behaviour of the UIP condition in EMDEs versus AEs in periods of uncertainty. The literature has shown that risk premia are higher for EMDEs (Aysun and Lee, 2013; Kumar, 2019); that the forward premium puzzle is mostly important for AEs (Bansal and Dahlquist, 2000; Frankel and Poonawala, 2010) and that global risk correlates positively with UIP deviations for both AEs and EMDEs, but that only for EMDEs policy uncertainty is a determinant of deviations (Kalemli-Özcan and Varela, 2021). More recently considering long-run UIP, Albagli et al. (2024) show that conditional on a global uncertainty shock, UIP violations increase both for emerging and advanced economies. With respect to these papers, our contribution is to do a comprehensive study of UIP deviations in EMDEs and AEs during periods of uncertainty looking both at the short- and long-run relationships.

<sup>&</sup>lt;sup>1</sup>Refer to Alper et al. (2009) for an excellent literature review on the UIP in EMDEs.

Section 2 presents the data used, section 3 presents he empirical estimations, section 4 the results and section 5 concludes.

#### 2 Data

We consider a sample of 60 countries for which data is available over the period 1995M1–2023M3. Among these countries, 39 are AEs (high-income and upper-middle-income countries), and 21 are EMDEs (lower-middle-income and low-income countries).<sup>2</sup>

We compute ex-post UIP deviations over a 3-month horizon considering the US as the reference country. We employ the 10-year government bond yields for testing long-run UIP.

We use different financial uncertainty indicators: VIX, World Uncertainty Index (WUI), Banking, currency, default and inflation composite index (BCDI), Financial Stress (FS), Financial Uncertainty (FU).

Finally, we employ macroeconomic variables as controls: inflation differentials, exchange rate regimes, capital controls. Appendix A presents the data in detail.

#### 3 Empirical analysis

We estimate the standard Fama regression conditioned on the income group:

$$s_{c,t} - s_{c,t-3} = \alpha + \beta_1 i dif f_{c,t-3} + \beta_2 \iota_c + \beta_3 \left( i dif f_{c,t-3} \times \iota_c \right) + \delta_c + \epsilon_t \tag{1}$$

We then estimate the same model adding uncertainty (VIX). Finally, we turn to a non-linear model estimation, in which the slope of the UIP is conditioned on both the uncertainty indicator and the income group:

$$s_{c,t} - s_{c,t-3} = \alpha + \beta_1 i dif f_{c,t-3} + \beta_2 V I X_{t-3} + \beta_3 (i dif f_{c,t-3} \times V I X_{t-3}) + \beta_4 (i dif f_{c,t-3} \times \iota_c \times V I X_{t-3})$$
(2)  
  $+ \beta_5 (i dif f_{c,t-3} \times \iota_c) + \beta_6 (\iota_c \times V I X_{t-3}) + \delta_c + \delta_{rg} + \delta_t + \epsilon_t$ 

As standard in this literature,  $s_{c,t} - s_{c,t-3}$  is the ex-post three-month change in the nominal exchange rate of country c;  $idiff_{c,t-3}$  is the difference between the country c nominal interest rate and the one of the US  $(i_{c,t-3} - i_{US,t-3})$ ; VIX is the proxy for global risk perception; and  $\iota$  is an income group dummy which takes the value 0 for low and lower-middle income countries and 1 otherwise.  $\delta$  represents fixed effects: country (c), regional-quinquennium  $(rq)^3$  and time (t). The constant is  $\alpha$  and the error term is represented by  $\epsilon_t$ .

 $<sup>^2</sup>$ A country is assigned to a constant World Bank income category, according to the group in which it most frequently falls under.

<sup>&</sup>lt;sup>3</sup>The region selection is based on the World Bank regions of 2022: East Asia and Pacific (1), Europe and Central Asia (2), Latin America and the Caribbean (3), Middle East and North Africa (4), North America (5), South Asia (6), Sub-Saharan Africa (7)).

(6) (1) (2)(3)(4)(5)0.109\*\*\* 0.108\*\*\* idiff0.028\*0.026\*0.070\*\* 0.083\*\* (0.033)(0.014)(0.014)(0.032)(0.033)(0.039)VIX0.073\*\*\* 0.012\*0.079\*\*\* (0.006)(0.016)(0.016)-0.003\*\*  $idiff \times VIX$ -0.004\*\*\* -0.003\*\* -0.004\*\* (0.001)(0.001)(0.001) $idiff \times VIX \times \iota$ 0.006\*\*\* 0.005\*\* 0.005\* $0.005^{\circ}$ (0.002)(0.002)(0.002)(0.002)-0.130\*\*\* -0.111  $idiff \mathbf{x} \iota$ -0.035\* -0.036\* -0.152\*-0.126\*\*(0.021)(0.021)(0.046)(0.047)(0.047)(0.053) $VIX \mathbf{x} \iota$ -0.080 -0.082-0.084-0.110\*\* (0.018)0.018 0.018 (0.023)No. Obs. 18,994 18,994 18,994 18.994 18,994 15,001 Adjusted R<sup>2</sup> 0.0717 0.0718 0.07260.0763 0.0883 0.0953 Fixed effects c, rq c, rq, t c, rq, t Other Controls No No No No No Yes

Table 1: Main regressions using VIX as the uncertainty indicator

In column (6) the macroeconomic control variables considered are inflation differentials, exchange rate regime, and capital controls. \*\*\*, \*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively.

#### 4 Results

#### 4.1 Main results

Table 1 shows the baseline results.<sup>4</sup> Column (1) presents the estimation of the model in equation 1; column (2) adds the VIX as a regressor; and columns (3)-(6) present different forms of the estimation of the model in equation 2.

Columns (1)-(2) show that the UIP is not satisfied, on average, over the whole countries' sample (coefficient of idiff different from unity) and that the VIX explains only partly the exchange rate volatility. Columns (3)-(6) display the key result. In periods of uncertainty, the deviations from the UIP are stronger for developing countries (negative coefficient of the interaction term  $idiff \times VIX$ ) whereas they narrow for advanced economies (positive coefficient of the triple interaction). Moreover, in periods of uncertainty, only advanced economies suffer from the so-called "forward premium puzzle" which is the tendency of a currency to appreciate when the country's interest rate is higher than the one of the reference country (the overall value of the coefficient idiff is negative for these economies). This evidence is robust to controlling for region-business cycle and time-fixed effects—column (5)— and to including macroeconomic variables—column (6)—as controls. In detail, we control for i) inflation differentials to gauge expected monetary policy conduct and real value of the currencies; ii)

<sup>&</sup>lt;sup>4</sup>Appendix B presents the results of the estimations at the country and income-group level. The results are coherent with Table 1.

exchange rate regime, and iii) capital controls to account for impediments to the free movement of capital as required by the UIP condition.

#### 4.2 Different measures of uncertainty

Table 2 confirms the previous results. In particular, it seems that the uncertainty measures for which the heterogeneous behaviour of UIP risk premiums across EMDEs and AEs is more marked are those linked to international financial uncertainty (BDCI, FS, FU). We do not find evidence of the same heterogeneous behaviour for other macroeconomic uncertainty indicators (see Appendix B). This reinforces our conjecture that in periods of financial uncertainty, international investors change their risk preferences and move their investments from EMDEs to AEs. These behaviours are consistent with Bhattarai et al. (2020) who shows that US uncertainty shock negatively affects EMEs exchange rates, raises EMEs country spreads, and reduces their capital inflows.

Table 2: Main regression using different financial uncertainty indicators

	VIX	WUI	BCDI	FS	$\mathbf{FU}$
idiff	0.108***	0.191***	0.130**	5.811***	0.127**
	(0.039)	(0.071)	(0.039)	(1.61)	(0.064)
$idiff \mathbf{x} uncert$	-0.004**	-0.000	-0.002***	-0.057***	-0.115*
	(0.001)	(0.000)	(0.000)	(0.015)	(0.068)
$idiff \mathbf{x} \ uncert \mathbf{x} \ \iota$	0.005**	-0.000	0.004***	0.075***	0.189**
	(0.002)	(0.000)	(0.001)	(0.023)	(0.092)
$idiff \mathbf{x} \iota$	-0.126**	-0.126	-0.211***	-7.666***	-0.182**
	(0.053)	(0.115)	(0.059)	(2.400)	(0.085)
$uncert \mathbf{x} \iota$	-0.110**	0.000	-0.064***	-0.696***	-3.933***
	(0.023)	(0.000)	(0.009)	(0.222)	(0.933)
No. Obs.	15,001	7,369	9,488	13,275	15,001
Adjusted R <sup>2</sup>	0.0953	0.0960	0.1026	0.1011	0.0951
Fixed effects	c, rq, t	c, rq, t	c,rq,t	c, rq, t	c, rq, t
Other Controls	Yes	Yes	Yes	Yes	Yes

In all regressions, control variables include inflation differentials, exchange rate regimes, capital controls. \*\*\*, \*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively.

#### 4.3 Main results over time

Table 3 shows the evolution of our main results over time, along two lines: by splitting the sample for short-run UIP and by considering long-run UIP.

We split the sample around the Global Financial Crisis (GFC) using two key dates: 2007m1 as in Engel et al. (2022) and 2008m8 following Andrews et al. (2024). In line with the previous literature, we observe a shift in the beta coefficient of the UIP relationship and, more related to our results, a stronger heterogeneous effect of VIX across countries in the second part of the sample.

This difference can be explained by the higher level of uncertainty in the recent period.

For long-run UIP, we regress the 10-year government bonds' yield differentials over the 10-year exchange rates' change in line with Chinn and Meredith (2004). We find evidence that our main results hold also for long-run UIP.

Before Long-run UIP From Baseline 2007 m12008m8 2007 m12008m810y-10y5.682\*\*\* idiff0.108\*\*\* 0.110\*\* 0.084\*0.208\*\*\* 0.195\*\*\* (0.039)(0.053)(0.050)(0.064)(0.068)(2.170) $idiff \ge VIX$ -0.004\*\* -0.002-0.002-0.002-0.003-0.119(0.001)(0.002)(0.002)(0.002)(0.089)(0.002) $idiff \times VIX \times \iota$ 0.005\*\*0.0020.0020.009\*\*0.008\*\* 0.537\*\*\* (0.002)(0.003)(0.002)(0.004)(0.004)(0.127)-0.434\*\*\* -0.370\*\*\*  $idiff \mathbf{x} \iota$ -29.128\* -0.126\*\*-0.081-0.057(0.116)(0.053)(0.069)(0.065)(0.105)(3.634) $VIX \times \iota$ -0.110\*\* -0.154\*\*\* -0.138\*\*\* -0.073\*\* -0.077\*\* -3.094\*\*\* (0.023)(0.039)(0.036)(0.032)(0.034)(1.125)No. Obs. 4,935 15,001 6,864 7,835 8,137 7,166 Adjusted R<sup>2</sup> 0.0953 0.1320 0.1139 0.07830.1019 0.5803 Fixed effects c, rq, t Other Controls Yes Yes Yes Yes Yes Yes

Table 3: Main regressions over time

In all regressions, control variables include inflation differentials, exchange rate regimes, capital controls. \*\*\*, \*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively.

#### 5 Conclusion

We study UIP deviations in periods of uncertainty by differentiating between advanced and developing-emerging countries. We find that UIP deviations widen in periods of uncertainty for developing and emerging countries, whereas they narrow for advanced economies. This evidence is stronger since the global financial crisis and finds validation also when considering long-run UIP. Our results hold across various financial uncertainty measures and do not seem to be triggered by macroeconomic uncertainty.

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#### A Data

Table A.1 presents the data used in the estimations and its sources. Table A.2 shows the correlation across different uncertainty indicators. Before running the econometric analysis, we preprocess the data and remove the outliers by constructing the excess returns' variable:

$$X_{c,t} = ediff_{c,t-3} - idiff_{c,t-3}$$
(A.1)

and dropping the 1st and 99th percentiles from its distribution.

For the long-run UIP regression, we follow the same procedure after constructing the excess returns variable corresponding to the data used.

Using 10-year government bond yields reduces the sample to 35 countries which are composed of 25 AEs and 10 EMDEs: Armenia Australia, Austria, Belgium, Canada, Czech Rep, Denmark, Egypt, France, Germany, Ghana, Hungary, Iceland, India, Italy, Japan, Kenya, Korea Rep, Malawi, Malaysia, Moldova, New Zealand, Nigeria, Norway, Philippines, Poland, Sierra Leone, Singapore, South Africa, Spain, Sweden, Tanzania.

Table A.1: Macroeconomic and uncertainty variables

Variable	Description	Temporal coverage	Source
Uncertainty indicators			
VIX	Measure of stock market's expectations of volatility based on S&P 500 $$	Monthly 1995M01 - 2023M03	Chicago Board Options Exchange
BCDI	Composite world index based on country-specific data of 70 countries on banking crisis, sovereign defaults, defaults on domestic debt, inflation crises and currency crises.	Yearly 1995 - 2010	Reinhart and Rogoff (2010)
USREC	It is based on the US Business cycle Expansions and Contractions data provided by NBER. The time series is composed of dummy variables when takes the value of 1 during a recessionary period in the US while takes 0 during an expansionary period.	Monthly 1995M01 - 2023M03	FRED Economic data
WUI	The World Uncertainty is a text-based crisis indicator that is computed by counting the word "uncertain" or similar words in the Economist Intelligence Unit country reports. The index measure uncertainty all over the world. It is the weighted average of 71 countries	Monthly 2008M01 - 2023M03	Ahir et al. (2022)
Monetary Policy Uncertainty (MPU)	MPU is a text-based indicator for the US that counts the words associated with monetary policy uncertainty across hundreds of daily newspapers covered by Acess World News in the US.	Monthly 1995M01 - 2023M03	Baker et al. (2016)
Trade Policy Uncertainty (TPU)	TPU is also a text-based indicator for the US based on news related to trade policy uncertainty in the US. The index reflects automated text-search results of the electronic archives of 7 leading newspapers discussing trade policy uncertainty: Boston Globe, Chicago Tribune, Guardian, Los Angeles Times, New York Times, Wall Street Journal, and Washington Post	Monthly 1995M01 - 2023M03	Caldara et al. (2020)
Financial Stress (FS)	Text-based indicator for the US based on titles of articles published in five US newspapers: the Boston Globe, Chicago Tribune, Los Angeles Times, Wall Street Journal and Washington Post.	Monthly 1995M01 - 2016M12	Püttmann (2018)
Financial Uncertainty (FU)	The aggregate of the volatility of statistical forecasts for a large number of financial series such as valuation ratios, growth rates of the aggregate dividends and prices, default and term spreads, yield spreads as well as risk factors like market risk premiums. US based indicator.	Monthly 1995M01 - 2022M12	Jurado et al. (2015)
Macroeconomic Uncertainty (MU)	The aggregate of the volatility of statistical forecasts for a large number of macroeconomic time series like real output and income, employment and hours, real retail, manufacturing and trade sales, consumer spending, etc. The indicator is constructed by the factor augmented vector autoregression. US based indicator.	Monthly 1995M01 - 2022M12	Jurado et al. (2015)
Real Uncertainty (RU)	Aggregates the volatility of statistical forecasts of the macroeconomic and financial series combined. US based indicator.	Monthly 1995M01 - 2022M12	Jurado et al. (2015)
Macroeconomic variables		Monthly	International financial statistics
Interest Rates (short-run)	3-months Interest Rates on Government Bonds, Securities and Treasury Bills $ \\$	1995M01 - 2023M03	(IMF) and DATASTREAM
Interest Rates (long-run)	10 year government bond yields.	Monthly 1995M01 - 2023M03	DATASTREAM
Exchange rates	Exchange rates (National Currency Per U.S. Dollar)	Monthly 1995M01 - 2023M03	International financial statistics (IMF) and DATASTREAM
Inflation differentials	We compute the inflation differential as the difference in the domestic CPI and US CPI. No data is available for Australia, Belize, New Zealand, and Vanuatu. Data starts from 2002 for Namibia, from 2006 for Lesotho and Sierra Leone	Monthly 1995M01 - 2023M03	Macroeconomic and Financial Data (IMF)
Exchange rate regime	Regime with the highest frequency for each country is taken for all the years till 2023M3. The dataset does not include Eswatini but since its currency has always been pegged to South African Rand, ER regime is given the value 1. No data available for Zambia since 2017	Monthly 1995M01 - 2019M12	Ilzetzki et al. (2019)
Capital controls	The capital control variable is a dummy variable that takes the value 0 for a unified market and 1 if there are dual, multiple, or parallel rates $$	Monthly 1995M01 - 2021M06	Ilzetzki et al. (2019)
Country classification	Country classification by income level The countries are assigned the income categories which they most frequently fall under. That income category is kept constant through the years for that country to ensure the results are not driven by sample composition effects.	Monthly 1995M01 - 2021M12	World bank data 2022
Region classification	Region classification by geographical zones. The countries are indexed according to the region that they are located in, which are - East Asia and Pacific, Europe and Central Asia, Latin America and the Carribean, Middle East and North Africa, North America, South Asia and Sub-Saharan Africa	Monthly 1995M01 - 2021M12	World bank data 2022

VIX BCDI USREC WUI MPUTPU $\mathbf{FS}$  $\mathbf{FU}$ MURUVIX BCDI0.1606USREC 0.44860.69031  $\mathbf{WUI}$ -0.1994-0.4756-0.5583MPU-0.0352 0.38830.29020.40271  $\mathbf{TPU}$ 0.03670.2130-0.0547 -0.1992 0.17250.8583 0.5189 -0.2289 -0.0399 FS0.35610.46111  $\mathbf{FU}$ 0.84750.22560.5854-0.33490.19540.01700.8114MU0.7566 0.3531 0.6840 -0.5243-0.03460.7979 0.9172 0.1150 RU0.74660.15660.5055-0.42950.0047-0.04970.77110.92780.9621

Table A.2: Correlation between different uncertainty indicators

Table A.3: Summary Statistics of Uncertainty Variables

				Percentiles				
	Mean	SD	5%	50%	95%	Min	Max	N
VIX	20.33	7.86	11.56	18.63	34.54	9.51	59.89	18994.00
USREC	0.08	0.28	0.00	0.00	1.00	0.00	1.00	18994.00
WUI	22547.05	9044.29	12265.10	20446.37	39762.47	9050.27	57517.98	10197.00
BCDI	35.84	22.86	0.68	36.45	72.25	0.68	72.25	10833.00
FS	101.19	0.90	99.87	101.15	102.59	99.54	105.89	14978.00
FU	0.92	0.19	0.68	0.89	1.22	0.63	1.55	18994.00
MPU	87.22	56.48	29.81	71.18	189.91	16.57	407.94	18994.00
TPU	45.30	40.06	21.15	31.69	141.71	11.30	266.00	18994.00
MU	0.66	0.12	0.54	0.62	0.91	0.53	1.22	18994.00
RU	0.64	0.13	0.54	0.60	0.88	0.53	1.39	18994.00

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### B Additional regressions

#### B.1 Country-by-country and income-group regressions

Table B.1: Income-group regressions

		Emerging	and Deve	loping Cou	intries	
	(1)	(2)	(3)	(4)	(5)	(6)
$\overline{idiff}$	0.028	0.024	0.109**	0.040	0.057	0.083
	(0.019)	(0.019)	(0.043)	(0.044)	(0.047)	(0.058)
VIX		0.039***	0.073***	0.083***	-	-
		(0.014)	(0.021)	(0.021)		
$idiff \times VIX$			-0.004**	-0.002	-0.003*	-0.004
			(0.001)	(0.001)	(0.002)	(0.002)
No. Obs.	6,342	6,342	6,342	6,342	6,342	4,795
Adjusted R <sup>2</sup>	0.1031	0.1040	0.1046	0.1193	0.1230	0.1320
Fixed effects	c	$\mathbf{c}$	c	c, rq	c, rq, t	c, rq, t
Other Controls	No	No	No	No	No	Yes
		Ac	dvanced E	conomies		
	(1)	(2)	(3)	(4)	(5)	(6)
$\overline{idiff}$	-0.007	-0.007	-0.043*	-0.040	-0.018	0.019
	(0.012)	(0.012)	(0.036)	(0.027)	(0.028)	(0.032)
VIX		-0.000	-0.006	-0.005	-	-
		(0.06)	(0.007)	(0.007)		
$idiff \times VIX$			0.001	0.002*	0.001	0.001
			(0.001)	(0.001)	(0.001)	(0.001)
No. Obs.	12,652	12,652	12,652	12,652	12,652	10,206
Adjusted $\mathbb{R}^2$	0.0062	0.0093	0.0063	0.0076	0.0237	0.0218
Fixed effects	c	$\mathbf{c}$	c	c, rq	c, rq, t	c, rq, t
Other Controls	No	No	No	No	No	Yes

In all regressions, control variables include inflation differentials, exchange rate regimes, capital controls. \*\*\*, \*\*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively. The separation into EMDEs and AEs follows the Income group definition in table B.2.

Table B.2: Country-by-country regressions

Country	(1)	(:	2)		(3)		Inc.	Inc.	Reg.
	$eta_1$	$eta_1$	$eta_4$	$eta_1$	$eta_2$	$\beta_3$	class	group	code
Antigua and Barb	0.000	0.000	0.000	-0.001	0.000	0.000	1	1	3
Australia	-0.006**		***0.002***		-0.003*		1	1	1 2 3 4 3 2 5
Austria	-0.086*		-0.008	-0.008	-0.003	-0.003	1	1	$\frac{2}{2}$
Bahamas	0.000 0.000***	0.000	0.000 **0.000*	0.000	$0.000 \\ 0.000*$	0.000	$\frac{1}{1}$	1	3
Bahrian Barbados	0.000	0.000	0.000	$0.000 \\ 0.000$	0.000	0.000	1	$\frac{1}{1}$	4 3
Belgium	0.007	0.078	-0.034	-0.319*		0.018*	i	i	2
Canada	0.000	0.001	-0.001**	0.006	-0.001*	0.000	î	ī	5
Hong Kong	-0.001	0.000	0.000*	-0.004	0.000*	0.000	i	i	ĭ
Czech Republic	0.060	0.069*	-0.009	0.048	-0.011	0.001	1	1	
Denmark 1	-0.011	-0.010	-0.002	0.007	0.000	-0.001	1	1	2
France	-0.038**		*-0.004	0.003	-0.001	-0.002	1	1	2
Germany	-0.009**		*-0.001	0.003	0.000	-0.001*	1	1	2 2 2 2 2 2 2 4 2
Greece	-0.007	-0.021	-0.104**	-0.153	-0.135*		1	1	2
Hungary	0.141*		*-0.036	0.027	-0.086	0.007	1	1	2
Iceland	0.064 0.006***	0.052	0.031 **0.001		**-*0.176* 0.000			$\frac{1}{1}$	2
Isreal Italy	1 /25***	• 0.005 ° • 1.600*	*0.001 *0.225***	-0.001 -3.977*	り.UUU *内 119	$0.000 \\ 0.134*$	** 1	1	4
Japan	-0.040	-0.010	-0.082*	-0.856	-0.025	0.134 $0.041$	1	1	$\frac{2}{1}$
Korea Rep.	-0.643	-0.660	0.029	-0.577	0.043	-0.005	1	1	1
Luxembourg	-0.353**		-0.023	-0.312	-0.024	0.000	ī	ī	$\overset{1}{2}$
New Zealand	-0.001	-0.001	-0.002***	* -0.024*	***0.004*		ī	$\bar{1}$	$\bar{1}$
Norway	-0.035**	-0.026	-0.005	-0.039	-0.005	0.001	Ī	$\bar{1}$	$\frac{1}{2}$
Polanď	0.004***	0.004*	**0.000	0.001	-0.001	0.000	1	1	$\frac{2}{1}$
Singapore	-0.002*	-0.002	0.000	0.012*		-0.001*		1	
Spain	-0.137	-0.144	0.037*	0.086	0.063	-0.012	1	1	2
Sweden Trinidad and Tob	-0.031**	-0.028*	-0.004	-0.072*	-0.005	0.002	1	1	2
United Kingdom	$0.000 \\ -0.727$	$0.000 \\ -0.707$	$0.000 \\ 0.331$	$-0.002 \\ 0.692$	$0.000 \\ 0.818$	$0.000 \\ -0.071$	$\frac{1}{1}$	$\frac{1}{1}$	<u>ქ</u>
Belize	0.000	0.000	0.000	0.000	0.000	0.000		1	2 2 3 2 3 3 1
Brazil	0.003*	U UU3*	-0.001	0.001	0.001	0.000	$\frac{2}{2}$	ī	$\check{3}$
Fiji	-0.008***	°-0.008*	*0.000	-0.016*	0.000	0.000	2	1	1
Jamaica	-0.034*	-0.043*	**0.067***	-0.088	0.038	0.002	2 2 2 ** 2	1	3 1 3 7 2 7 7
Malaysia	0.010***	0.010*	**0.000	0.029** 0.002		-0.001*	2	$\frac{1}{1}$	1
Mexico Namibia	0.000 -0.036**	0.000	-0.009 -0.005	-0.002	-0.008 $0.004$	$0.000 \\ -0.001$	2	1	3 7
Romania	0.002***	0.002*	**D.000	-0.003	-0.004	0.000*	** 2	1	2
Sevchelles	-0.035*	-0 044*	*0.024	-0.096	0.011	0.002	$\bar{2}$	$\bar{1}$	$\overline{7}$
South Africa	-0.045***	· -0.032*	-0.013*	-0.019	-0.009	-0.001	2	1	7
Algeria	0.306***	$-0.297^*$	^10.007	0.355*	0.009	-0.003	2 3 3 3	0	$^4$
Armenia	1.096***	0.928*	**0.269	1.589*	* 0.536*	-0.033	3	0	$\begin{smallmatrix}2\\7\\4\end{smallmatrix}$
Cabo Verde	$-0.084 \\ 0.028*$	-0.067	-0.018 * 0.007	$0.089 \\ 0.004$	-0.001 -0.006	-0.008 $0.001$	ა ე	$_{0}^{0}$	1
Egypt Eswatini	-0.054**		**-0.003	-0.023	0.006	-0.001	3	0	7
Guyana	0.060	0.041	0.048**	-0.492*	-0.055	0.034 *	$\frac{3}{3}$	ŏ	$\frac{7}{3}$
India	0.208***	0.208*	* <u>*</u> 0.002	0.040	-0.050	0.009	$\ddot{3}$	ŏ	6
Lesotho	-0.048***	`-0.045*	*-0.003	0.015	0.016	-0.003	3	ŏ	$\begin{array}{c} 6 \\ 7 \\ 2 \\ 7 \end{array}$
Moldova	0.013**		* 0.012*	-0.057*	**0.030*		** 3	Ō	$\underline{2}$
Nigeria	-0.182	-0.137	0.297***		0.346*	-0.007	3	0	
Philippines	0.078 0.011***	0.081*	-0.006 * 0.002**	0.129	0.002	-0.002	3	0	1
Solomon Islands Vanuatu	-0.310**	0.010	*-0.02	-0.012 $0.454$	-0.001 $0.212$	0.001* -0.037	3	0	1
Ghana	-0.310		**0.003**	0.454	**0.212 **0.012**	**0.03 <i>1</i>	** 4	0	1 7 7 7 6 7 7
Kenya	-0.038	-0.032	-0.035	0.373*	*0.131 *	*-0.020*	$**\overset{\tau}{4}$	ŏ	ż
Malawi	0.017	0.007	-0.186**	-0.039	-0.224	0.002	$\bar{4}$	Ŏ	7
Nepal	0.130***	0.126 *		0.361*		-0.011*	4	0	6
Sierra Leone	-0.010***			-0.005	0.005	0.000	4	0	7
Tanzania	0.220	0.319*	0.283*	0.839* -0.011*	0.536*	-0.033*	4	0	$\frac{7}{2}$
Uganda Zambia	-0.002 -0.007*	-0.002 -0.008*	0.000 $0.007$	-0.011	0.000 $-0.011$	$0.000 \\ 0.001$	$\frac{4}{4}$	0 0	7
Zambia	-0.007	-0.008	0.007	-0.022	-0.011	0.001	4	- 0	

Column (1) refers to the regression's coefficient in equation 1; Column (2) refers to the regression's coefficients in equation 1 with the addition of the VIX ( $\beta_4$ ); and Column (3) refers to the regression's coefficients in equation 2. "Inc. class" refers to the World Bank income classification, while "Inc. group" refers to the income group dummy used in the panel estimations. "Reg. code" is the World Bank regional classification. \*\*\*, \*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively.

#### B.2 Macroeconomic uncertainty indicators

Table B.3 shows results macroeconomic uncertainty indicators like the US Business Cycle (USREC), Trade Policy Uncertainty (TPU), Monetary Policy Uncertainty (MPU), Macroeconomic Uncertainty (MU), Real Uncertainty (RU).

Table B.3: Main regression using macroeconomic uncertainty indicators

	USREC	TPU	MPU	MU	RU
idiff	0.030	0.060**	0.030	-0.092	-0.200
	(0.019)	(0.026)	(0.025)	(0.102)	(0.158)
$idiff \mathbf{x} uncert$	-0.007	-0.000*	-0.000	0.200	0.398
	(0.041)	(0.000)	(0.000)	(0.164)	(0.269)
$idiff \mathbf{x} uncert \mathbf{x} \iota$	0.045	0.001	0.000	-0.018	-0.131
	(0.062)	(0.000)	(0.000)	(0.224)	(0.360)
$idiff \mathbf{x} \iota$	-0.026	-0.059	-0.0354	-0.010	0.052
	(0.026)	(0.038)	(0.036)	(0.139)	(0.212)
$uncert \mathbf{x} \iota$	-0.090	0.001	-0.000	-2.186	-5.128
	(0.627)	(0.004)	(0.000)	(2.176)	(3.39)
No. Obs.	15,001	15,001	15,001	15,001	15,001
Adjusted $\mathbb{R}^2$	0.0939	0.0943	0.0939	0.0943	0.0947
Fixed effects	c, rq, t				
Other Controls	Yes	Yes	Yes	Yes	Yes

In all regressions, control variables include inflation differentials, exchange rate regimes, capital controls. \*\*\*, \*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively.

#### B.3 Long-run UIP

Table B.4 shows the results for long-run UIP using the 10-year sovereign bond yields as an explanatory variable and, as explained variable either i) the 3months exchange rate differential as in the baseline (column "3m - 10y")<sup>5</sup>; and ii) the 10-year exchange rates' difference (column "10y - 10y") as done in Chinn and Meredith (2004).

 $<sup>^5</sup>$ This procedure is coherent with the study of Albagli et al. (2024) which compares returns to short horizon investments in long-term bonds.

Table B.4: Long-run UIP regressions

	3m - 10y	10y - 10y
$\overline{idiff}$	0.088	5.682***
	(0.134)	(2.170)
$idiff \times VIX$	-0.002	-0.119
	(0.006)	(0.089)
$idiff \times VIX \times \iota$	0.002	0.537***
	(0.009)	(0.127)
$idiff \mathbf{x} \iota$	0.208	-29.128***
	(0.255)	(3.634)
$VIX \mathbf{x} \iota$	-0.092	-3.094***
	(0.082)	(1.125)
No. Obs.	$7,\!360$	4,935
Adjusted $\mathbb{R}^2$	0.1460	0.5803
Fixed effects	c, rq, t	c, rq, t
Other Controls	Yes	Yes

In all regressions, control variables include inflation differentials, exchange rate regimes, capital controls. \*\*\*, \*\*, and \* denote statistical significance at the 1, 5, and 10% levels, respectively.